# Distribution Fitting 4. Benford test on a sample of observed genotypes number from running of a genetic algorithm 

Lorentz JÄNTSCHI ${ }^{1)}$, Sorana D. BOLBOACĂ ${ }^{2}$, , (armen E. STOENOIU ${ }^{1)}$, Mihaela IANCU ${ }^{2)}$, Monica M. MARTA ${ }^{2)}$, Elena M. PICĂ ${ }^{(1)}$, Monica ŞTEFU ${ }^{1)}$, Adriana F. SESTRAŞ ${ }^{3)}$, Marcel M. DUDA ${ }^{3}$, Radu E. SESTRAŞ ${ }^{3)}$, Ştefan ȚIGAN ${ }^{2)}$, Ioan ABRUDAN $^{1)}$, Mugur C. BĂLAN ${ }^{1)}$<br>${ }^{1)}$ Technical University of Cluj-Napoca, 103-105 Muncii Bvd., 400641 Cluj, Romania<br>${ }^{2)}$ Iuliu Hațieganu University of Medicine and Pharmacy, Cluj-Napoca, 400349 Cluj, Romania<br>${ }^{3)}$ University of Agricultural Sciences and Veterinary Medicine of Cluj-Napoca, 3-5 Mănăştur St., 400372, Cluj, Romania<br>correspondence: lori@,academicdirect.org


#### Abstract

A new designed genetic algorithm was run on experimental data obtained from measurements of octanol-water partition coefficients of a series of polychlorinated biphenils, in order to relate their structure with their activity. A family of molecular descriptors having all necessary ingredients to run a genetic algorithm on it characterized the structure. An experiment using different selection and survival strategies were conducted, when multiple runs were recorded and their results were analyzed. Total number and number of distinct of genotypes present in the generations leading to evolution were included in an analysis concerning the validity of data with Benford test, the methodology and the obtained results being given.


Keywords: distribution fitting; statistical agreement; genotypes number; genetic algorithm; survival strategy; selection strategy; Benford statistic

## INTRODUCTION

The Benford test uses the normal distribution to check if for an array of numbers their digits follow the Benford distribution.

The hypothesis of the test is that the values of the observations measurements are often logarithmically distributed and thus the logarithm of the measurement set is uniform distributed. The distribution, test, and its statistic are called after the physician Frank BENFORD, who discovered first and it formulated intuitively (Benford, 1938), inspired in its survey by a short communication of Simon NEWCOMB (Newcomb, 1881). The proof of the distribution it comes later being given by Theodore P. HILL (Hill, 1995).

This intuitively result of counting digits occurrences of the numbers was found true to a large variety of datasets, including electricity bills (Christian and others, 1993), forensic and financial audits (de Marchi and Hamilton, 2006; Nigrini and Mittermaier, 1997), stock exchanges (Ley, 1996), river basin area, weights of chemicals and streets addresses (Benford, 1938), roundoff errors (Barlow and Bareiss 1985), population numbers (Sandron, 2002), ceasing rates (Leemis and others, 2000), physical and mathematical constants and a lot of processes described by power lows common in nature (Newcomb, 1881; Berger and Hill, 2007).

An important result is that the result (once observed when the number is expressed in a numeration base) is independent of the numeration base in which the numbers were expressed, even if the proportions of representation are changing (Pinkham, 1961; Hill, 1995).

A natural consequence of the existence of the Benford law is that this fact can be used to validate the reported data under the presumption of the altering (mystification) of them, the immediate approach being the comparing of the observed first digit frequencies with the theoretical ones (Diekmann, 2007; Günnel and Tödter, 2008).

The Benford test were run on the results giving numbers of alive genotypes obtained from runs of a genetic algorithm searching on structure-activity relationships between the structure of a series of 206 biphenyl polychlorinated compounds and their observed octanolwater partition coefficient, in order to test if these numbers follows the Bendford law.

## MATERIALS AND METHODS

In forty-six independent runs for every combination of survival method and selection method from Deterministic, Tournament and Proportional types were recorded the evolutions of a genetic algorithm (Jäntschi, 2009). On these data, for twenty equal width classes of observation, from generation 1 to the generation 20000, the total number of genotypes and the number of distinct genotypes were grouped and the observed frequencies were obtained (Table 1).

Table 1. Observed frequencies of the viable genotypes in the generations leading to evolution from forty-six independent runs for different selection and survival strategy

| SelSrv | Millennium | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13 | 14 | 15 | 16 | 17 | 18 | 19 | 20 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| PP | num_obs | 732 | 141 | 97 | 62 | 36 | 44 | 25 | 30 | 35 | 32 | 29 | 16 | 20 | 15 | 8 | 7 | 12 | 21 | 17 | 0 |
| PP | sum_obs | 8366 | 1580 | 1073 | 700 | 417 | 500 | 285 | 339 | 390 | 358 | 330 | 186 | 221 | 166 | 92 | 84 | 134 | 245 | 201 | 235 |
| PT | num_obs | 713 | 171 | 133 | 89 | 54 | 44 | 40 | 33 | 32 | 32 | 11 | 28 | 32 | 26 | 10 | 21 | 6 | 8 | 1 | 15 |
| PT | sum obs | 8207 | 1960 | 1505 | 1013 | 609 | 505 | 455 | 383 | 363 | 369 | 127 | 313 | 362 | 279 | 111 | 239 | 69 | 90 | 128 | 161 |
| PD | num_obs | 748 | 159 | 101 | 85 | 60 | 71 | 24 | 25 | 28 | 21 | 22 | 33 | 24 | 25 | 10 | 15 | 15 | 22 | 14 | 18 |
| PD | sum_obs | 8825 | 1858 | 1166 | 990 | 702 | 829 | 273 | 286 | 326 | 243 | 257 | 38 | 278 | 292 | 116 | 171 | 176 | 256 | 162 | 207 |
| TP | num_obs | 655 | 159 | 110 | 88 | 57 | 31 | 37 | 33 | 34 | 27 | 25 | 23 | 17 | 9 | 7 | 8 | 9 | 12 | 10 | 3 |
| TP | sum_obs | 7470 | 1741 | 1202 | 992 | 639 | 331 | 410 | 363 | 372 | 300 | 278 | 262 | 191 | 100 | 79 | 92 | 99 | 134 | 115 | 147 |
| TT | num_obs | 740 | 178 | 92 | 62 | 56 | 37 | 38 | 31 | 21 | 23 | 19 | 25 | 14 | 17 | 18 | 25 | 9 |  | 7 | 12 |
| TT | sum_obs | 8475 | 1988 | 1025 | 700 | 625 | 424 | 427 | 352 | 237 | 248 | 218 | 269 | 159 | 195 | 199 | 278 | 102 | 12 | 82 | 136 |
| TD | num_obs | 757 | 110 | 98 | 83 | 61 | 66 | 49 | 39 | 21 | 31 | 29 | 25 | 31 | 16 | 5 | 4 | 14 | 9 | 11 | 6 |
| TD | sum_obs | 8969 | 1282 | 1145 | 963 | 704 | 769 | 569 | 453 | 244 | 364 | 333 | 291 | 364 | 187 | 59 | 48 | 165 | 105 | 129 |  |
| DP | num_obs | 422 | 94 | 44 | 33 | 43 | 28 | 15 | 43 | 31 | 15 | 15 | 17 | 12 | 17 | 16 | 9 | 15 | 17 | 6 | 5 |
| DP | sum_obs | 4739 | 987 | 450 | 349 | 467 | 289 | 150 | 451 | 344 | 167 | 166 | 183 | 116 | 163 | 157 | 99 | 150 | 190 | 68 | 57 |
| DT | num_obs | 431 | 110 | 65 | 51 | 48 | 39 | 38 | 45 | 36 | 43 | 14 | 25 | 29 | 13 | 8 | 16 | 8 | 10 | 11 | 2 |
| DT | sum_obs | 4883 | 1223 | 719 | 558 | 533 | 440 | 411 | 477 | 385 | 468 | 153 | 273 | 312 | 139 | 92 | 169 | 87 | 111 | 119 | 20 |
| DD | num_obs | 466 | 120 | 81 | 66 | 53 | 41 | 24 | 39 | 51 | 24 | 11 | 19 | 18 | 23 | 23 | 14 | 19 | 12 | 24 | 2 |
| DD | sum obs | 5511 | 1402 | 949 | 772 | 627 | 486 | 283 | 459 | 596 | 28 | 130 | 225 | 212 | 271 | 270 | 166 | 222 | 14 | 285 | 24 |

SelSrv: 46 runs using Sel and Srv as selection and survival strategies;
Srv, $\operatorname{Sel} \in\{\mathrm{P}, \mathrm{T}, \mathrm{D}\} ; \mathrm{P}$ - Proportional strategy; T - Tournament strategy; D - Deterministic strategy;
Millennium = 1000 generations; num_obs: number of distinct genotypes; sum_obs: total number of genotypes;
First three digits of the numbers from Table 1 were included into the analysis of the agreement with Benford distribution.

The theoretical probabilities of the Benford distribution for first ( $\mathrm{d}_{0}$, Eq. 1 ), second $\left(\mathrm{d}_{1}\right.$, Eq. 2 ) and third ( $\mathrm{d}_{2}$, Eq. 3 ) digits are given by following relationships:

$$
\begin{align*}
& p\left(d_{0}\right)=\log _{b}\left(1+1 / d_{0}\right), d_{0}=1 . .(b-1)  \tag{1}\\
& p\left(d_{1}\right)=\sum_{k=1}^{b-1} \log _{b}\left(1+1 /\left(k \cdot b+d_{1}\right)\right), d_{1}=0 . .(b-1) \tag{2}
\end{align*}
$$

$$
\begin{equation*}
\mathrm{p}\left(\mathrm{~d}_{2}\right)=\sum_{\mathrm{j}=1}^{\mathrm{b}-1} \sum_{\mathrm{k}=0}^{\mathrm{b}-1} \log _{\mathrm{b}}\left(1+1 /\left(\mathrm{j} \cdot \mathrm{~b}^{2}+\mathrm{k} \cdot \mathrm{~b}+\mathrm{d}_{2}\right)\right), \mathrm{d}_{2}=0 . .(\mathrm{b}-1) \tag{3}
\end{equation*}
$$

where $b$ is the numeration base ( $b=10$ for the data given in Table 1 ).

## RESULTS AND DISCUSSION

The theoretical (Benford) and the observed relative differences are given in Figure 1.




| $\operatorname{CDF}\left(\mathrm{d}_{1}\right)$ |
| :--- | :--- |
| $\square$ Benford $\square$ Difference |





Benford: Theoretical distribution; Difference: observed difference relative to theoretical
Figure 1. Probability distribution functions (PDF) and cumulative distribution functions (CDF) for first three digits ( $\mathrm{d}_{0}, \mathrm{~d}_{1}$, and $\mathrm{d}_{2}$ respectively) of the data from Table 1

The plots from Figure 1 show a small disagreement between observation and the model of the Benford distribution. In order to measure its observing probability, two statistics were involved: Chi Square (Table 2) and Kolmogorov-Smirnov (Table 3); their results are analyzed in the next.

Benford distribution does not have unknown (to be estimated) parameters. The numeration base (Eq.1-3) is known being 10, and thus the degrees of freedom are number of digits minus one (Table 2).

Table 2. $\chi^{2}$ test on observed frequencies of the data from Table 1 (null hypothesis: first, second, and third digits follows Benford distribution)

| Digit (i) |  |  | Expected frequency ( $\mathrm{E}_{\mathrm{i}}$ ) |  |  | Observed frequency ( $\mathrm{O}_{\mathrm{i}}$ ) |  |  | $\left(\left\|\mathrm{O}_{\mathrm{i}}-\mathrm{E}_{\mathrm{i}}\right\|-0.5\right)^{2}$ |  |  | $\left(\left\|\mathrm{O}_{\mathrm{i}}-\mathrm{E}_{\mathrm{i}}\right\|-0.5\right)^{2} / \mathrm{E}_{\mathrm{i}}$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\mathrm{d}_{0}$ | $\mathrm{d}_{1}$ | $\mathrm{d}_{2}$ | $\mathrm{d}_{0}$ | $\mathrm{d}_{1}$ | $\mathrm{d}_{2}$ | $\mathrm{d}_{0}$ | $\mathrm{d}_{1}$ | $\mathrm{d}_{2}$ | $\mathrm{d}_{0}$ | $\mathrm{d}_{1}$ | $\mathrm{d}_{2}$ | $\mathrm{d}_{0}$ | $\mathrm{d}_{1}$ | $\mathrm{d}_{2}$ |
| 0 | 0 | 0 | - | 40 | 19 | - | 28 | 25 | - | 144 | 36 | - | 3.31 | 1.59 |
| 1 | 1 | 1 | 108 | 38 | 19 | 117 | 41 | 18 | 81 | 9 | 1 | 0.67 | 0.16 | 0.01 |
| 2 | 2 | 2 | 63 | 37 | 18 | 72 | 37 | 18 | 81 | 0 | 0 | 1.15 | 0.01 | 0.01 |
| 3 | 3 | 3 | 45 | 35 | 18 | 48 | 33 | 20 | 9 | 4 | 4 | 0.14 | 0.06 | 0.13 |
| 4 | 4 | 4 | 35 | 34 | 18 | 33 | 34 | 11 | 4 | 0 | 49 | 0.06 | 0.01 | 2.35 |
| 5 | 5 | 5 | 29 | 33 | 18 | 17 | 42 | 15 | 144 | 81 | 9 | 4.56 | 2.19 | 0.35 |
| 6 | 6 | 6 | 24 | 32 | 18 | 16 | 34 | 18 | 64 | 4 | 0 | 2.34 | 0.07 | 0.01 |
| 7 | 7 | 7 | 21 | 31 | 18 | 18 | 30 | 18 | 9 | 1 | 0 | 0.30 | 0.01 | 0.01 |
| 8 | 8 | 8 | 18 | 30 | 18 | 19 | 31 | 16 | 1 | 1 | 4 | 0.01 | 0.01 | 0.13 |
| 9 | 9 | 9 | 17 | 29 | 18 | 20 | 29 | 23 | 9 | 0 | 25 | 0.37 | 0.01 | 1.13 |
| $\Sigma(\mathrm{df}=8)$ | $\Sigma(\mathrm{df}=9)$ | $\Sigma(\mathrm{df}=9)$ | 360 | 339 | 182 | 360 | 339 | 182 | 402 | 244 | 128 | 9.60 | 2.53 | 4.12 |
| $\mathrm{X}^{2}\left(\mathrm{~d}_{0}\right)=9.6 ; \mathrm{p} \chi^{2}(9.6,8)=29.4 \% ; \mathrm{X}^{2}\left(\mathrm{~d}_{1}\right)=2.53 ; \mathrm{p} \chi^{2}(2.53,9)=98.0 \% ; \mathrm{X}^{2}\left(\mathrm{~d}_{2}\right)=4.12 ; \mathrm{p} \chi^{2}(4.12,9)=90.3 \%$; |  |  |  |  |  |  |  |  |  |  |  |  |  |  |

Table 3. Kolmogorov-Smirnov test on observed cumulative frequencies of the data from Table 1 (null hypothesis: first, second, and third digits follows Benford distribution)

| Digit |  |  | Expected ( $\mathrm{d}_{0} \mathrm{e}, \mathrm{d}_{1} \mathrm{e}, \mathrm{d}_{2} \mathrm{e}$ ) and Observed ( $\left.\mathrm{d}_{0} \mathrm{O}, \mathrm{d}_{1} \mathrm{O}, \mathrm{d}_{2} \mathrm{o}\right)$ |  |  |  |  |  | Difference |  |  | Difference |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\mathrm{d}_{0}$ | $\mathrm{d}_{1}$ | $\mathrm{d}_{2}$ | $\mathrm{d}_{0} \mathrm{a}$ | $\mathrm{d}_{1} \mathrm{a}$ | $\mathrm{d}_{2} \mathrm{a}$ | $\mathrm{d}_{0} \mathrm{O}$ | $\mathrm{d}_{1} \mathrm{O}$ | $\mathrm{d}_{2} \mathrm{O}$ | $\mathrm{d}_{0}$ | $\mathrm{d}_{1}$ | $\mathrm{d}_{2}$ | $\mathrm{d}_{0}$ | $\mathrm{d}_{1}$ | $\mathrm{d}_{2}$ |
| 0 | 0 | 0 | 0 | 40 | 19 | 0 | 28 | 25 | 0 | 12 | -6 | 0 | 12 | 6 |
| 1 | 1 | 1 | 108 | 78 | 38 | 117 | 69 | 43 | -9 | 9 | -5 | 9 | 9 | 5 |
| 2 | 2 | 2 | 171 | 115 | 56 | 189 | 106 | 61 | -18 | 9 | -5 | 18 | 9 | 5 |
| 3 | 3 | 3 | 216 | 150 | 74 | 237 | 139 | 81 | -21 | 11 | -7 | 21 | 11 | 7 |
| 4 | 4 | 4 | 251 | 184 | 92 | 270 | 173 | 92 | -19 | 11 | 0 | 19 | 11 | 0 |
| 5 | 5 | 5 | 280 | 217 | 110 | 287 | 215 | 107 | -7 | 2 | 3 | 7 | 2 | 3 |
| 6 | 6 | 6 | 304 | 249 | 128 | 303 | 249 | 125 | 1 | 0 | 3 | 1 | 0 | 3 |
| 7 | 7 | 7 | 325 | 280 | 146 | 321 | 279 | 143 | 4 | 1 | 3 | 4 | 1 | 3 |
| 8 | 8 | 8 | 343 | 310 | 164 | 340 | 310 | 159 | 3 | 0 | 5 | 3 | 0 | 5 |
| 9 | 9 | 9 | 360 | 339 | 182 | 360 | 339 | 182 | 0 | 0 | 0 | 0 | 0 | 0 |
| $\Sigma$ | $\Sigma$ | $\Sigma$ | - | - | - | - | - | - | -66 | 55 | -9 | 82 | 55 | 37 |

$\mathrm{D}\left(\mathrm{d}_{0}\right)=21 ; \mathrm{K}\left(\mathrm{d}_{0}\right)=21 \sqrt{ } 9 / 360 ; \mathrm{pKS}(9,21 \sqrt{ } 9 / 360)=90.8 \%$;
$\mathrm{D}\left(\mathrm{d}_{1}\right)=12 ; \mathrm{K}\left(\mathrm{d}_{1}\right)=12 \sqrt{ } 10 / 339 ; \mathrm{pKS}(10,12 \sqrt{ } 10 / 339)=95.2 \%$;
$\mathrm{D}\left(\mathrm{d}_{1}\right)=7 ; \mathrm{K}\left(\mathrm{d}_{1}\right)=7 \sqrt{ } 10 / 182 ; \mathrm{pKS}(10,7 \sqrt{ } 10 / 182)=94.6 \%$;
High observation probabilities results from Chi Square test. A geometric mean of $64 \%$ of all three probabilities gives $64 \%$ probability to observe a worst agreement between the observed data following the Benford distribution and the theoretical Benford distribution under Chi-Square test hypothesis. Thus the hypothesis of Benford distribution of the digits from Table 1 cannot be rejected by using Chi Square statistic. Even higher probabilities results from Kolmogorov-Smirnov test. A geometric mean of $94 \%$ of all three probabilities gives $94 \%$ probability to observe a worst agreement between the observed data following the Benford distribution and the theoretical Benford distribution under Kolmogorov-Smirnov test
hypothesis. Thus the hypothesis of Benford distribution of the digits from Table 1 cannot be rejected by using Kolmogorov-Smirnov statistic.

The only one remained question about the analysis of the data from table 1 is regarding the difference between observation probabilities given by Chi Square test (64\%) and Kolmogorov-Smirnov test ( $94 \%$ ).

It's well known () that the Chi Square test assumes a normality distribution of the errors (squared in Table 2, $\left(\left|\mathrm{O}_{\mathrm{i}}-\mathrm{E}_{\mathrm{i}}\right|-0.5\right)^{2}$ column). A measure of the departure from normality may be given by the Jarque-Bera statistic (Jarque and Bera, 1981), as a measure of the sufficiency (Fisher, 1922) for Chi Square and Kolmogorov-Smirnov statistics. Under assumption of the Gauss distribution of the differences (when Chi Square has highest accuracy) the expected population kurtosis is 3, and is 6 under assumption of Laplace distribution of the differences (skewness expectation being zero). The kurtosis, the skewness and the Jarque-Bera statistic under two assumptions (Gauss and Laplace) for the differences from Table 2 are given in Table 4.

Table 4. Kurtosis and skewness of the disagreement between observed and the model

| Disagreement | $\mathrm{d}_{0}$ | $\mathrm{~d}_{1}$ | $\mathrm{~d}_{2}$ |
| :---: | :---: | :---: | :---: |
| Kurtosis | 2.13 | 4.60 | 2.71 |
| Skewness | 0.28 | 0.81 | 0.07 |
| Jarque-Bera(Gauss) | $0.41 ; \mathrm{p} \chi^{2}(0.41,2)=82 \%$ | $2.16 ; \mathrm{p} \chi^{2}(2.16,2)=34 \%$ | $0.04 ; \mathrm{p} \chi^{2}(0.04,2)=98 \%$ |
| Jarque-Bera(Laplace) | $25.1 ; p \chi^{2}(25.1,2)=10^{-6}$ | $4.33 ; \mathrm{p} \chi^{2}(4.33,2)=11 \%$ | $18.1 ; \mathrm{p} \chi^{2}(18.1,2)=10^{-4}$ |

Table 4 shows that the largest departure between the results obtained from Chi Square statistic and from Kolmogorov-Smirnov statistic is expected to be at $d_{0}$, followed by $d_{2}$ and the lowest difference should be at $d_{1}$.

Indeed, comparing the probabilities given in Table 2 with the ones given in Table 3, largest departure is at $\mathrm{d}_{0}(29 \%$ from Chi Square, Table 2; $90.8 \%$ from Kolmogorov-Smirnov, Table 3; difference: $61.4 \%$ ), followed by $\mathrm{d}_{2}(90.3 \%$ from Chi Square, Table 2; $94.6 \%$ from Kolmogorov-Smirnov, Table 3; difference: $4.3 \%$ ) and by $\mathrm{d}_{1}$ ( $98.0 \%$ from Chi Square, Table 2; $95.2 \%$ from Kolmogorov-Smirnov, Table 3; difference: 2.8\%).

## CONCLUSIONS

With a good confidence, given by the probability to observe the observed departures from Benford distribution, the numbers giving the total number and the number of distinct genotypes from independent runs of a genetic algorithm by using different selection and survival strategies follows Benford law.

Since (and when in general) the kurtosis of the differences between observation and the model shows closeness to the Gauss than the Laplace distribution, a more weight to the Chi Square statistic should be assigned when remarks regarding the probability of observation are made, and vice versa.

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## REFERENCES

Barlow, J. L. and E. H. Bareiss (1985). On Roundoff Error Distributions in Floating Point and Logarithmic Arithmetic. Computing 34:325-347.

Benford, F. (1938). The law of anomalous numbers. Proceedings of the American Philosophical Society 78(4):551-572.

Berger, A. and T. P. Hill (2007). Newton's method obeys Benford's law. American Mathematical Monthly 114(7):588-601.

Christian, C. W., S. Gupta and S. M. Lin (1993), Determinants of Tax Preparer Usage - Evidence From Panel-Data. National Tax Journal 46(4):487-503.
de Marchi, S. and J. T. Hamilton (2006). Assessing the accuracy of selfreported data: An evaluation of the toxics release inventory. Journal of Risk and Uncertainty 32(1):57-76.

Diekmann, A. (2007). Not the First Digit! Using Benford's Law to Detect Fraudulent Scientific Data. Journal of Applied Statistics. 34(3):321-329.

Fisher, R. A. (1920). A Mathematical Examination of the Methods of Determining the Accuracy of an Observation by the Mean Error, and by the Mean Square Error. Monthly Notices of the Royal Astronomical Society 80(S):758-770.

Fisher, R. A. (1922). On the Mathematical Foundations of Theoretical Statistics. Philosophical Transactions of the Royal Society A 222:309-368.

Günnel S. and K.-H. Tödter (2008). Does Benford's Law hold in economic research and forecasting? Empirica, DOI: 10.1007/s10663-008-9084-1.

Hill, T. P. (1995). Base invariance implies Benford's Law. Proceedings of the American Mathematical Society 123(3):887-895.

Hill, T. P. (1995). Base-Invariance Implies Benford's Law. Proceedings of the American Mathematical Society 123(3):887-895.

Jäntschi, L. (2009). A genetic algorithm for structure-activity relationships: software implementation. Manuscript: http://arxiv.org/abs/0906.4846 (abstract), http://arxiv.org/pdf/0906.4846 (PDF).

Jarque, C. M. and A. K. Bera. (1981). Efficient tests for normality, homoscedasticity and serial independence of regression residuals: Monte Carlo evidence. Economics Letters 7(4):313-318.

Leemis, L. M., B. W. Schmeiser and D. L. Evans (2000). Survival distributions satisfying Benford's law. The American Statistician 54(4):236-241.

Ley, E. (1996). On the Peculiar Distribution of the U.S. Stock Indices Digits. American Statistician 50:311-313.

Newcomb, S. (1881). Note on the Frequency of the Use of Digits in Natural Numbers. American Journal of Mathematics 4:39-40.

Nigrini, M. J. and L. J. Mittermaier (1997). The use of Benford's Law as an aid in analytical procedures. Auditing-A Journal of Practice \& Theory 16(2):52-67.

Pinkham, R. S. (1961). On the distribution of first significant digits. Annals of Mathematical Statistics 32(4):1223-1230.

Sandron, F. (2002). Do Populations Conform to the Law of Anomalous Numbers?. Population 57(4-5):755-762.

